Machine Learning CSE 6363 (Fall 2016)

Lecture 7 Fisher Linear Discriminant Analysis

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The figure on the right shows greater separation between subsets, one set of the points with dashed line, another with solid line.

- Suppose we have 2 classes and *d*-dimensional samples *x*₁,...,*x*_n where
 - n₁ samples come from the first class
 - n_2 samples come from the second class
- consider projection on a line
- Let the line direction be given by unit vector v



- Scalar **v**^t**x**_i is the distance of projection of **x**_i from the origin
- Thus it **v**^t**x**_i is the projection of **x**_i into a one dimensional subspace

- Thus the projection of sample x_i onto a line in direction v is given by v^tx_i
- How to measure separation between projections of different classes?
- Let $\tilde{\mu}_1$ and $\tilde{\mu}_2$ be the means of projections of classes 1 and 2
- Let μ_1 and μ_2 be the means of classes 1 and 2
- $|\tilde{\mu}_1 \tilde{\mu}_2|$ seems like a good measure

$$\widetilde{\mu}_1 = \frac{1}{n_1} \sum_{x_i \in C_1}^{n_1} \boldsymbol{v}^t \boldsymbol{x}_i = \boldsymbol{v}^t \left(\frac{1}{n_1} \sum_{x_i \in C_1}^{n_1} \boldsymbol{x}_i \right) = \boldsymbol{v}^t \boldsymbol{\mu}_1$$

similarly, $\tilde{\mu}_2 = \mathbf{v}^t \mu_2$

- How good is $|\tilde{\mu}_1 \tilde{\mu}_2|$ as a measure of separation?
 - The larger $|\tilde{\mu}_1 \tilde{\mu}_2|$, the better is the expected separation



- the vertical axes is a better line than the horizontal axes to project to for class separability
- however $\left| \hat{\mu}_1 \hat{\mu}_2 \right| > \left| \hat{\mu}_1 \hat{\mu}_2 \right|$

• The problem with $|\tilde{\mu}_1 - \tilde{\mu}_2|$ is that it does not consider the variance of the classes







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- We need to normalize $|\tilde{\mu}_1 \tilde{\mu}_2|$ by a factor which is proportional to variance
- Have samples z_1, \dots, z_n . Sample mean is $\mu_z = \frac{1}{n} \sum_{i=1}^n z_i$
- Define their *scatter* as $\boldsymbol{s} = \sum_{i=1}^{n} (\boldsymbol{z}_i - \boldsymbol{\mu}_z)^2$
- Thus scatter is just sample variance multiplied by n
 - scatter measures the same thing as variance, the spread of data around the mean
 - scatter is just on different scale than variance



- Fisher Solution: normalize $|\tilde{\mu}_1 \tilde{\mu}_2|$ by scatter
- Let $y_i = v^t x_i$, i.e. y_i 's are the projected samples
- Scatter for projected samples of class 1 is $\widetilde{\mathbf{S}}_{1}^{2} = \sum_{\mathbf{y}_{i} \in Class \ 1} (\mathbf{y}_{i} - \widetilde{\mu}_{1})^{2}$
- Scatter for projected samples of class 2 is $\widetilde{\boldsymbol{s}}_{2}^{2} = \sum_{\boldsymbol{y}_{i} \in Class \ 2} (\boldsymbol{y}_{i} - \widetilde{\boldsymbol{\mu}}_{2})^{2}$

- We need to normalize by both scatter of class 1 and scatter of class 2
- Thus Fisher linear discriminant is to project on line in the direction v which maximizes

want projected means are far from each other

$$\boldsymbol{J}(\boldsymbol{v}) = \frac{\left(\boldsymbol{\tilde{\mu}}_1 - \boldsymbol{\tilde{\mu}}_2\right)^2}{\boldsymbol{\tilde{S}}_1^2 + \boldsymbol{\tilde{S}}_2^2}$$

want scatter in class 1 is as small as possible, i.e. samples of class 1 cluster around the projected mean $\tilde{\mu}_1$ want scatter in class 2 is as small as possible, i.e. samples of class 2 cluster around the projected mean $\hat{\mu}_2$

$$\boldsymbol{J}(\boldsymbol{v}) = \frac{(\boldsymbol{\tilde{\mu}}_1 - \boldsymbol{\tilde{\mu}}_2)^2}{\boldsymbol{\tilde{S}}_1^2 + \boldsymbol{\tilde{S}}_2^2}$$

If we find v which makes J(v) large, we are guaranteed that the classes are well separated



$$\boldsymbol{J}(\boldsymbol{v}) = \frac{\left(\boldsymbol{\tilde{\mu}}_1 - \boldsymbol{\tilde{\mu}}_2\right)^2}{\boldsymbol{\tilde{S}}_1^2 + \boldsymbol{\tilde{S}}_2^2}$$

- All we need to do now is to express J explicitly as a function of v and maximize it
 - straightforward but need linear algebra and Calculus
- Define the separate class scatter matrices S₁ and S₂ for classes 1 and 2. These measure the scatter of original samples x_i (before projection)

$$\boldsymbol{S}_{1} = \sum_{\boldsymbol{x}_{i} \in Class \ 1} (\boldsymbol{x}_{i} - \boldsymbol{\mu}_{1}) (\boldsymbol{x}_{i} - \boldsymbol{\mu}_{1})^{t}$$
$$\boldsymbol{S}_{2} = \sum_{\boldsymbol{x}_{i} \in Class \ 2} (\boldsymbol{x}_{i} - \boldsymbol{\mu}_{2}) (\boldsymbol{x}_{i} - \boldsymbol{\mu}_{2})^{t}$$

Now define the *within* the class scatter matrix $S_w = S_1 + S_2$

• Recall that
$$\widetilde{\boldsymbol{s}}_{1}^{2} = \sum_{\boldsymbol{y}_{i} \in Class \ 1} (\boldsymbol{y}_{i} - \widetilde{\boldsymbol{\mu}}_{1})^{2}$$

• Using $\mathbf{y}_i = \mathbf{v}^t \mathbf{x}_i$ and $\hat{\mu}_1 = \mathbf{v}^t \mu_1$

$$\widetilde{\mathbf{S}}_{1}^{2} = \sum_{\substack{\mathbf{y}_{i} \in Class \ 1}} \left(\mathbf{v}^{t} \mathbf{x}_{i} - \mathbf{v}^{t} \boldsymbol{\mu}_{1} \right)^{2}$$

$$= \sum_{\substack{\mathbf{y}_{i} \in Class \ 1}} \left(\mathbf{v}^{t} (\mathbf{x}_{i} - \boldsymbol{\mu}_{1}) \right)^{t} \left(\mathbf{v}^{t} (\mathbf{x}_{i} - \boldsymbol{\mu}_{1}) \right)$$

$$= \sum_{\substack{\mathbf{y}_{i} \in Class \ 1}} \left((\mathbf{x}_{i} - \boldsymbol{\mu}_{1})^{t} \mathbf{v} \right)^{t} \left((\mathbf{x}_{i} - \boldsymbol{\mu}_{1})^{t} \mathbf{v} \right)$$

$$= \sum_{\substack{\mathbf{y}_{i} \in Class \ 1}} \mathbf{v}^{t} (\mathbf{x}_{i} - \boldsymbol{\mu}_{1}) (\mathbf{x}_{i} - \boldsymbol{\mu}_{1})^{t} \mathbf{v} = \mathbf{v}^{t} \mathbf{S}_{1} \mathbf{v}$$

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- Similarly $\tilde{\boldsymbol{s}}_{2}^{2} = \boldsymbol{v}^{t}\boldsymbol{S}_{2}\boldsymbol{v}$
- Therefore $\tilde{\boldsymbol{S}}_1^2 + \tilde{\boldsymbol{S}}_2^2 = \boldsymbol{v}^t \boldsymbol{S}_1 \boldsymbol{v} + \boldsymbol{v}^t \boldsymbol{S}_2 \boldsymbol{v} = \boldsymbol{v}^t \boldsymbol{S}_W \boldsymbol{v}$
- Define between the class scatter matrix $\boldsymbol{S}_B = (\mu_1 - \mu_2)(\mu_1 - \mu_2)^t$
- S_B measures separation between the means of two classes (before projection)
- Let's rewrite the separations of the projected means $(\tilde{\mu}_1 - \tilde{\mu}_2)^2 = (\mathbf{v}^t \mu_1 - \mathbf{v}^t \mu_2)^2$ $= \mathbf{v}^t (\mu_1 - \mu_2)(\mu_1 - \mu_2)^t \mathbf{v}$ $= \mathbf{v}^t \mathbf{S}_B \mathbf{v}$

Thus our objective function can be written:

$$\boldsymbol{J}(\boldsymbol{v}) = \frac{\left(\widetilde{\boldsymbol{\mu}}_{1} - \widetilde{\boldsymbol{\mu}}_{2}\right)^{2}}{\widetilde{\boldsymbol{s}}_{1}^{2} + \widetilde{\boldsymbol{s}}_{2}^{2}} = \frac{\boldsymbol{v}^{t}\boldsymbol{S}_{B}\boldsymbol{v}}{\boldsymbol{v}^{t}\boldsymbol{S}_{W}\boldsymbol{v}}$$

Maximize J(v) by taking the derivative w.r.t. v and setting it to 0

$$\frac{d}{dv}J(v) = \frac{\left(\frac{d}{dv}v^{t}S_{B}v\right)v^{t}S_{W}v - \left(\frac{d}{dv}v^{t}S_{W}v\right)v^{t}S_{B}v}{\left(v^{t}S_{W}v\right)^{2}}$$
$$= \frac{\left(2S_{B}v\right)v^{t}S_{W}v - \left(2S_{W}v\right)v^{t}S_{B}v}{\left(v^{t}S_{W}v\right)^{2}} = 0$$

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Need to solve $\mathbf{v}^{t}\mathbf{S}_{w}\mathbf{v}(\mathbf{S}_{B}\mathbf{v}) - \mathbf{v}^{t}\mathbf{S}_{B}\mathbf{v}(\mathbf{S}_{w}\mathbf{v}) = \mathbf{0}$ $\Rightarrow \frac{v^{t} S_{W} v(S_{B} v)}{v^{t} S_{W} v} - \frac{v^{t} S_{B} v(S_{W} v)}{v^{t} S_{W} v} = 0$ $\Rightarrow S_{B}V - \frac{V^{t}S_{B}V(S_{W}V)}{V^{t}S_{W}V} = 0$ $\Rightarrow \underbrace{\mathbf{S}_{B}\mathbf{v}}_{B} = \lambda \mathbf{S}_{W} \mathbf{v}$ generalized eigenvalue problem

$$\boldsymbol{S}_{B}\boldsymbol{v}=\lambda\boldsymbol{S}_{W}\boldsymbol{v}$$

 If S_W has full rank (the inverse exists), can convert this to a standard eigenvalue problem

$$\boldsymbol{S}_{W}^{-1}\boldsymbol{S}_{B}\boldsymbol{v}=\lambda\boldsymbol{v}$$

But $S_B x$ for any vector x, points in the same direction as $\mu_1 - \mu_2$

$$\boldsymbol{S}_{B}\boldsymbol{x} = (\mu_{1} - \mu_{2})(\mu_{1} - \mu_{2})^{t}\boldsymbol{x} = (\mu_{1} - \mu_{2})(\mu_{1} - \mu_{2})^{t}\boldsymbol{x} = \alpha(\mu_{1} - \mu_{2})^{t}\boldsymbol{x}$$

Thus can solve the eigenvalue problem immediately $v = S_w^{-1}(\mu_1 - \mu_2)$

$$\mathbf{S}_{W}^{-1}\mathbf{S}_{B}[\mathbf{S}_{W}^{-1}(\mu_{1}-\mu_{2})] = \mathbf{S}_{W}^{-1}[\alpha(\mu_{1}-\mu_{2})] = \alpha[\mathbf{S}_{W}^{-1}(\mu_{1}-\mu_{2})]$$

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Fisher Linear Discriminant Example

Data

- Class 1 has 5 samples c₁=[(1,2),(2,3),(3,3),(4,5),(5,5)]
- Class 2 has 6 samples c₂=[(1,0),(2,1),(3,1),(3,2),(5,3),(6,5)]
- Arrange data in 2 separate matrices

$$\boldsymbol{c}_1 = \begin{bmatrix} \boldsymbol{1} & \boldsymbol{2} \\ \vdots & \vdots \\ \boldsymbol{5} & \boldsymbol{5} \end{bmatrix} \qquad \boldsymbol{c}_2 = \begin{bmatrix} \boldsymbol{1} & \boldsymbol{0} \\ \vdots & \vdots \\ \boldsymbol{6} & \boldsymbol{5} \end{bmatrix}$$

 Notice that PCA performs very poorly on this data because the direction of largest variance is not helpful for classification



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Fisher Linear Discriminant Example

- First compute the mean for each class $\mu_1 = mean (c_1) = \begin{bmatrix} 3 & 3.6 \end{bmatrix}$ $\mu_2 = mean (c_2) = \begin{bmatrix} 3.3 & 2 \end{bmatrix}$
- Compute scatter matrices S_1 and S_2 for each class $S_1 = 4 * cov(c_1) = \begin{bmatrix} 10 & 8.0 \\ 8.0 & 7.2 \end{bmatrix}$ $S_2 = 5 * cov(c_2) = \begin{bmatrix} 17.3 & 16 \\ 16 & 16 \end{bmatrix}$
- Within the class scatter:

$$\mathbf{S}_{W} = \mathbf{S}_{1} + \mathbf{S}_{2} = \begin{bmatrix} 27.3 & 24 \\ 24 & 23.2 \end{bmatrix}$$

it has full rank, don't have to solve for eigenvalues

- The inverse of S_W is $S_W^{-1} = inv(S_W) = \begin{bmatrix} 0.39 & -0.41 \\ -0.41 & 0.47 \end{bmatrix}$
- Finally, the optimal line direction \mathbf{v} $\mathbf{v} = \mathbf{S}_{w}^{-1}(\mu_{1} - \mu_{2}) = \begin{bmatrix} -0.79\\ 0.89 \end{bmatrix}$

Fisher Linear Discriminant Example

- Notice, as long as the line has the right direction, its exact position does not matter
- Last step is to compute the actual *1D* vector *y*.
 Let's do it separately for each class



$$Y_{1} = v^{t}c_{1}^{t} = \begin{bmatrix} -0.65 & 0.73 \end{bmatrix} \begin{bmatrix} 1 \cdots 5 \\ 2 \cdots 5 \end{bmatrix} = \begin{bmatrix} 0.81 \cdots 0.4 \end{bmatrix}$$
$$Y_{2} = v^{t}c_{2}^{t} = \begin{bmatrix} -0.65 & 0.73 \end{bmatrix} \begin{bmatrix} 1 \cdots 6 \\ 0 \cdots 5 \end{bmatrix} = \begin{bmatrix} -0.65 \cdots -0.25 \end{bmatrix}$$

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Within-class Covariance Matrix

- The projection is from a d-dimensional space to a (c-1) dimensional space.
- The within-class scatter is

$$S_w = \sum_{i=1}^c S_i$$

$$S_i = \sum_{X \in D_i} (X - m_i) (X - m_i)^t$$

$$m_i = \frac{1}{n_i} \sum_{X \in D_i} X$$

Total Covariance Matrix

• Total mean

$$m = \frac{1}{n} \sum_{X} X$$

• Total scatter matrix

$$S_{T} = \sum_{X}^{c} (X - m)(X - m)^{t}$$

$$S_{T} = \sum_{i=1}^{c} \sum_{X \in D_{i}}^{c} (X - m_{i} + m_{i} - m)^{t}$$

$$(X - m_{i} + m_{i} - m)^{t}$$

$$S_{T} = \sum_{i=1}^{c} \sum_{X \in D_{i}}^{c} (X - m_{i})(X - m_{i})^{t}$$

+

$$\sum_{i=1}^{c} \sum_{X \in D_{i}} (m_{i} - m)(m_{i} - m)^{t}$$

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Machine Learning 21

Total Covariance Matrix

• Note

$$\sum_{i=1}^{c} \sum_{X \in D_{i}} (X - m_{i})(m_{i} - m)^{t} =$$

$$\sum_{i=1}^{c} \sum_{X \in D_{i}} (X - m_{i})[(m_{i} - m)^{t}] = [0]$$

$$0 - matrix$$

$$S_{T} = S_{W} + \sum_{i=1}^{c} n_{i}(m_{i} - m)(m_{i} - m)^{t}$$

$$Define$$

$$S_{B} = \sum_{i=1}^{c} n_{i}(m_{i} - m)(m_{i} - m)^{t}$$

$$S_T = S_W + S_B$$

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Multiple Discriminant Analysis

• The projection from a d-dimensional space to a (c-1)dimensional space is done by c-1 discriminant functions

 $y_i = w_i^t X \quad i = 1, \dots, c$

 y_i can be viewed as a component of a vector Y

- W_i are viewed as columns of a d X (c-1) matrix $Y = W^t X$
- The samples $X_1, ..., X_n$ (d-dimensional) are mapped to a set of $y_1, ..., y_n$ (c-1)-dimensional which can be described by their own mean vectors and scatter matrices

Multiple Discriminant Analysis



It can be shown

$$\widetilde{S}_W = W^t S_W W$$
$$\widetilde{S}_B = W^t S_B W$$

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Multiple Discriminant Analysis

• We want to find a transform matrix W that maximizes the ratio of the determinants of the between-class scatter to the within-class scatter:

$$J(W) = \frac{\left|\tilde{S}_{B}\right|}{\left|\tilde{S}_{W}\right|} = \frac{\left|W^{t}S_{B}W\right|}{\left|W^{t}S_{W}W\right|}$$

• The columns of an optimal W are the generalized eigenvectors corresponding to the largest eigenvalues

$$S_B w_i = \lambda_i S_W w_i$$

• If S_W is nonsingular, $S_W^{-1}S_B W_i = \lambda_i W_i$

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FDA and MDA Drawbacks

- Reduces dimension only to k = c-1 (unlike PCA)
 - For complex data, projection to even the best line may result in unseparable projected samples
- Will fail:
 - **1.** J(v) is always 0: happens if $\mu_1 = \mu_2$



 If J(v) is always large: classes have large overlap when projected to any line (PCA will also fail)

