#### **Generalized Additive Models**

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 The Linear Regression Model and its Smooth Extension
 The Generalized Linear Models
 Smooth Extensions of Generalized Linear Models and the Training Method for the Generalized Additive Models

#### Linear Regression

1. The Linear Regression Model and its Smooth Extension

Response Variable: Y
 A set of predictor random variables X1, X2, ..., Xp
 A set of n independent realizations

 (y1, x11, ..., x1p), ..., (yn, xn1, ..., xnp)

**Regression Procedure:** 

$$E(Y | X_1, X_2, \dots, X_p)$$
  
=  $\beta_0 + \beta_1 X_1 + \dots + \beta_p X_p$ 

Given a sample, estimate coefficients

### Smooth Extension(Additive Model)

The additive Model:

$$E(Y | X_1, X_2, \dots, X_p) = s_0 + \sum_{j=1}^p s_j(X_j)_j$$

Sj(.) are smooth functions

Estimate coefficients





• Our Model:

$$E(Y \mid X) = s(X)$$

- Any reasonable estimate of E(Y | X = x)
- Local average estimates (Ni: neighborhood of xi)

 $\hat{s}(x_i) = \operatorname{Ave}_{j \in N_i} \{y_j\}$ 

$$N_{i} = \left\{ \max\left(i - \frac{[wn] - 1}{2}, 1\right), \dots, i - 1, i, \\ i + 1, \dots, \min\left(i + \frac{[wn] - 1}{2}, n\right) \right\}$$

#### Scatterplot Smoothers

Any reasonable estimate of E(Y | X = x)

 $\hat{s}(x_i) = \operatorname{Ave}_{j \in N_i} \{y_j\}$ 

- Local average estimates (Ni: neighborhood of xi)
  - 1) Running mean ---> arithmetic mean
  - 2) Running lines smoother

$$\hat{\beta}_{0i}=\bar{y}_i-\hat{\beta}_{1i}\bar{x}_i$$

$$\hat{s}(x_{i}) = \hat{\beta}_{0i} + \hat{\beta}_{1i}x_{i}$$

$$\hat{\beta}_{1i} = \frac{\sum_{j \in N_{i}} (x_{j} - \bar{x}_{i})y_{j}}{\sum_{j \in N_{i}} (x_{j} - \bar{x}_{i})^{2}}$$

where: 
$$\bar{x}_i = (1/n) \sum_{j \in N_i} x_j$$
  $\bar{y}_i = (1/n) \sum_{j \in N_i} y_j$ 

#### Span Selection(Window size)

**Consider the Extreme Choices** 

- W = 1/n Estimate of S(xi) is yi high Variance, not smooth
- W = 2 Estimate of S(xi) is global least squares regression too smooth, might be biased
  - Window Size(Span) is chosen between 1/n and 2
    - ----> Bias-Variance Tradeoff
  - Minimize Cross-Validation Sum of Squares

$$\text{CVSS}(w) = (1/n) \sum_{1}^{n} (y_i - \hat{s}_w^{-i}(x_i))^2.$$



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- General linear model
   Y = b0 + b1\*X1 + ... + bm\*Xm
- Generalized linear model
   Y = g(b0 + b1\*X1 + ... + bm\*Xm)
- Formally, the inverse function of g(...), gi(...) is called the link function gi(muY) = b0 + b1\*X1 + ... + bm\*Xm muY: expected value of Y

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gi(muY) = b0 + b1*X1 + ... + bm*Xm
muY: expected value of Y
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Algorithm for estimate of b=(b0, b1, ..., bm):

- Adjusted dependent variable regression Given:
   Etah = b0 + b1\*X1 + ... + bm\*Xm (current estimate of linear predictor) muY: corresponding fitted value
- Adjusted dependent variable:

$$z = \hat{\eta} + (y - \hat{\mu}) \left( \frac{d\eta}{d\mu} \right)$$

Adjusted dependent variable:

$$z = \hat{\eta} + (y - \hat{\mu}) \left( \frac{d\eta}{d\mu} \right)$$

• Define weights W:

$$(W)^{-1} = \left(\frac{d\eta}{d\mu}\right)^2 V,$$

V: the variance of Y at mu = muh

Adjusted dependent variable regression algorithm:

$$z = \hat{\eta} + (y - \hat{\mu}) \left( \frac{d\eta}{d\mu} \right)$$

- 1) Algorithm regressiong z on x1, ..., xp with weights W,
- 2) Obtain an estimate coefficient b,
- 3) Using estimate b to computer the estimate of mu and eta,
- 4) A new z is computed
- 5) Repeat 1) to 4) until the change in deviance is small enough:

$$\operatorname{dev}(y,\,\hat{\mu}) = 2[l(y) - l(\hat{\mu})]$$



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#### Smooth Extension of GLM

• S1(.), ..., Sp(.) are smooth functions:

$$\eta = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p$$

$$\eta = s_0 + \sum_{1}^{p} s_j(X_j)$$
Generalized Additive Model:
$$g(\mu) = \eta$$

### Estimation – Backfitting Algorithm

Suppose 
$$Y = s_0 + \sum_{j=1}^p s_j(X_j) + \varepsilon$$
 is correct

Define the partial residual:

$$R_j = Y - s_0 - \sum_{k \neq j} s_k(X_k)$$

Then:

$$E(R_j \mid X_j) = s_j(X_j)$$

Minimize:

$$E\left(Y-s_0-\sum_{j=1}^p s_j^m(X_j)\right)^2$$

#### Estimation – Backfitting Algorithm

Backfitting Algorithm Initialization:  $s_0 = E(Y), \quad s_1^1(\cdot) \equiv s_2^1(\cdot) \equiv \cdots$  $\equiv s_p^1(\cdot) \equiv 0, \quad m = 0.$ Iterate: m = m + 1for j = 1 to p do:  $R_j = Y - s_0 - \sum_{k=1}^{j-1} s_k^m(X_k)$  $-\sum_{k=1}^{p}s_{k}^{n-1}(X_{k})$  $s_i^m(X_j) = E(R_j \mid X_j).$ Until: RSS =  $E\left(Y - s_0 - \sum_{j=1}^p s_j^m(X_j)\right)^2$ fails to decrease.

#### General Local Scoring Algorithm

General Local Scoring Algorithm Initialization:  $s_0 = g(E(y)), \quad s_1^0(\cdot) \equiv s_2^0(\cdot) \equiv \cdots$  $\equiv s_p^0(\cdot) \equiv 0, \quad m = 0.$ 

Iterate: m = m + 1

1. Form the adjusted dependent variable

$$Z = \eta^{m-1} + (Y - \mu^{m-1})(\partial \eta / \partial \mu^{m-1}),$$

where

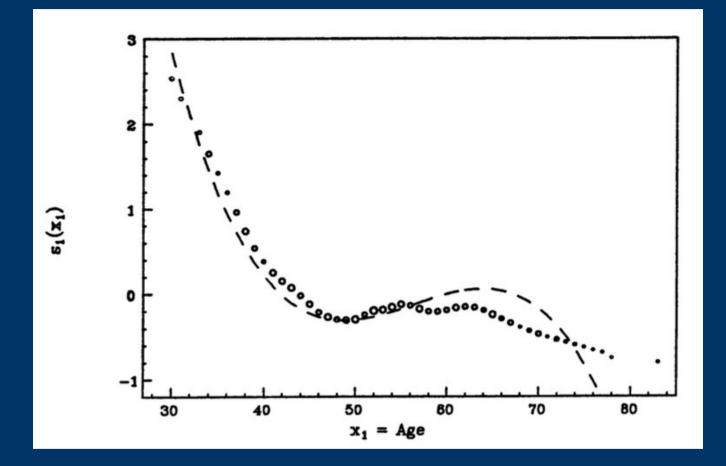
$$\eta^{m-1} = s_0 + \sum_{j=1}^p s_j^{m-1}(X_j)$$
 and

$$\eta^{m-1}=g(\mu^{m-1}).$$

- 2. Form the weights  $W = (\partial \mu / \partial \eta^{m-1})^2 V^{-1}$ .
- 3. Fit an additive model to Z using the backfitting algorithm with weights W, we get estimated functions  $s_j^m(\cdot)$  and model  $\eta^m$ .

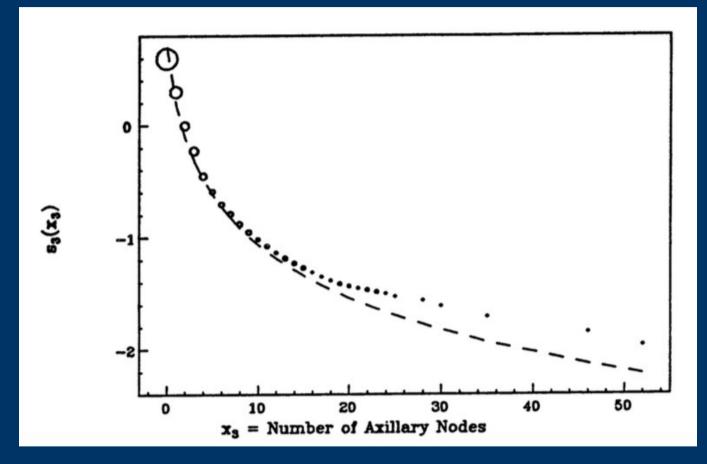
Until:  $E \operatorname{dev}(Y, \mu^m)$  fails to decrease.

# Experiment



Breast Cancer Dataset

# Experiment



Breast Cancer Dataset

# Thanks!